A Randomized Controlled Trial of Mindfulness-Based Cognitive Therapy for Treatment-Resistant Depression

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Key Words
Cognitive therapy · Depression · Major depressive disorder · Mindfulness meditation · Mindfulness-based psychotherapy · Treatment-resistant depression

Abstract
Background: Due to the clinical challenges of treatment-resistant depression (TRD), we evaluated the efficacy of mindfulness-based cognitive therapy (MBCT) relative to a structurally equivalent active comparison condition as adjuncts to treatment-as-usual (TAU) pharmacotherapy in TRD. Methods: This single-site, randomized controlled trial compared 8-week courses of MBCT and the Health Enhancement Program (HEP), comprising physical fitness, music therapy and nutritional education, as adjuncts to TAU pharmacotherapy for outpatient adults with TRD. The primary outcome was change in depression severity, measured by percent reduction in the total score on the 17-item Hamilton Depression Rating Scale (HAM-D17), with secondary depression indicators of treatment response and remission. Results: We enrolled 173 adults; mean length of a current depressive episode was 6.8 years (SD = 8.9). At the end of 8 weeks of treatment, a multivariate analysis showed that relative to the HEP condition, the MBCT condition was associated with a significantly greater mean percent reduction in the HAM-D17 (36.6 vs. 25.3%; p = 0.01) and a significantly higher rate of treatment responders (30.3 vs. 15.3%; p = 0.03). Although numerically superior for MBCT than for HEP, the rates of remission did not significantly differ between treatments (22.4 vs. 13.9%; p = 0.15). In these models, state anxiety, perceived stress and the presence of personality disorder had adverse effects on outcomes. Conclusions: MBCT significantly decreased depression severity and improved treatment response rates at 8 weeks but not remission rates. MBCT appears to be a viable adjunct in the management of TRD.

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Introduction

Major depressive disorder is a debilitating disease estimated to affect up to 14 million adults annually in the United States [1]. Despite many pharmacological treatment options for depression, as many as 60% of patients treated with treatment-as-usual (TAU) pharmacotherapy fail to fully remit after 1 year [2]. Clinical consensus regarding the definition of treatment-resistant depression (TRD) is wide-ranging following in the footsteps of Thase and Rush [3], who developed the first staging pro-
tocol for treatment resistance. Currently, many experts define TRD as the failure to remit after at least two antidepressant trials of adequate dose and duration [4]. Current TRD definitions are somewhat limited since they generally do not include psychotherapy trials in their classification systems. Many individuals with TRD will suffer chronic depression, typically diagnosed when depressive symptoms have been present for 2 years or more. Although in DSM IV-TR chronic major depression and less symptomatic dysthymia were differentiated, in DSM V, they have been combined into persistent depressive disorder [5]. We should note that TRD is sometimes incorrectly identified (i.e., pseudo treatment resistance), when there has been inadequate diagnosis or treatment [6]. Even with multiple medication trials, however, treatment resistance is quite common [7].

Although most studies of TRD have focused on medication strategies, recent approaches have utilized evidence-based psychotherapeutic interventions. In the STAR*D study, patients at level 2 who had not responded adequately to an initial trial with citalopram were randomized to cognitive behavioral therapy (CBT) or a change in medication [8]. CBT was offered either alone or as augmentation to continued citalopram. Medication changes were either switches or augmentations. All of the arms had similar remission rates of approximately 25%, although the medication augmentation strategies achieved significantly faster remission compared to CBT augmentation [9].

In an earlier study by Paykel et al. [10], CBT was compared to medication management in patients with incomplete remission. Both groups showed similar reductions in depression ratings, although CBT appeared to offer protection against relapse. The cognitive behavioral analysis system of psychotherapy (CBASP) was investigated in comparison to nefazodone in patients with chronic depression, although only a minority were treatment resistant. Both medication and CBASP produced individual remission rates of approximately 30% and had a rate of 48% when combined [11]. Kocsis et al. [12] investigated individuals with chronic depression, including 491 patients with TRD. The study compared CBASP or supportive psychotherapy augmentation of medication treatment with a flexible, optimized course of pharmacotherapy alone. In this study, neither form of psychotherapy was superior to the pharmacotherapy alone. In 2015, Brakemeier et al. [13] investigated the application of CBASP in an inpatient setting with 70 patients with chronic depression and treatment resistance. In this uncontrolled pre-post design, 40% of patients remitted after the 12-week intervention. At 52 weeks, 48% of patients had a sustained response. In uncontrolled studies of mindfulness-based cognitive therapy (MBCT) for TRD, both Kenny and Williams [14] in a study of 79 patients and Eisendrath et al. [15] in a study of 55 patients demonstrated large effect sizes (>1.0) in reducing depression severity.

TRD is associated with poorer clinical outcome and adherence [16], as well as higher rates of disability, mortality and morbidity, and lower overall quality of life compared to those with nonresistant unipolar depression [17, 18]. Given these factors, there is an urgent need for innovative and effective treatment strategies for TRD.

Because of this need and our pilot data [15], we examined the possibility that MBCT might be effective as an adjunctive treatment for TRD. Studies have suggested that MBCT strengthens cognitive control of emotion regulation [19]. TRD is believed to be related to a decreased regulation of the dorsal executive control system of the ventral affective processing system [20–22]. Strengthening the dorsal system (e.g., dorsolateral prefrontal cortex or dorsal anterior cingulate) through meditation effects such as enhanced selective attention, executive functioning and emotion regulation helps restore cognitive control processes to down-regulate the ventral system (such as amygdala and ventrolateral prefrontal cortex) and depression-related processes such as rumination [23, 24]. As Barnhofer et al. [25, 26] have pointed out, by strengthening cognitive control mechanisms, meditation may lessen ruminative processes and thereby diminish depression. This theoretical background led us to hypothesize that MBCT could prove effective in diminishing TRD. We specifically hypothesized that MBCT would be more effective than another credible intervention that controlled for nonspecific effects.

MBCT is an 8-week group program originally developed as an intervention to protect against relapse for patients with unipolar depression who had achieved remission after initial treatment [19]. Multiple randomized controlled trials (RCTs) have since confirmed its effectiveness in preventing relapse when compared to routine clinical management or maintenance antidepressant treatment [19, 27–29]. Because most of the RCTs evaluating MBCT for relapse prevention used treatment as usual as the control condition, it was difficult to ascertain whether sustained treatment effects were related to specific components of MBCT or nonspecific effects.

With this limitation in mind, we designed this study to evaluate the effects of MBCT as an adjunct to TAU pharmacotherapy for adults with TRD compared to a
structurally equivalent comparator condition plus TAU. The comparator, the Health Enhancement Program (HEP), controlled for the following nonspecific factors: group support and morale; reduction in stigma, facilitator attention; treatment duration, and time spent at home practice. Because of these HEP features, we believed the two arms of the study offered clinical equipoise deserving investigation, especially since MBCT had never been studied in an RCT in the TRD population at study initiation. We named the study ‘Practicing Alternative Techniques to Heal Depression’ to avoid biasing participants with investigators’ hypotheses. Secondary aims included exploration of potential mediators of MBCT effects, including mindfulness, self-compassion, rumination and experiential avoidance. We predicted that MBCT would be superior to HEP primarily in reducing depression severity while also enhancing response and remission rates.

Methods

Study Design
We conducted the study at the University of California, San Francisco (UCSF), with a 4-year period of active recruitment from September 2009 to September 2013. The study design and rationale have previously been published [30]. The primary outcome analysis evaluated percent change in depression severity after an 8-week acute treatment, the primary endpoint of the study. Clinical evaluations were conducted in person by research assistants, who held advanced degrees in psychology and were blind to treatment assignment. To assure blinding, participants were instructed not to discuss treatment assignment with research assistants prior to assessments, and both intervention groups were held on the same day and time to minimize raters having cues to participant assignment. Lastly, treatment groups were held in the evenings, outside of traditional work hours, when the research assistants had already left the research offices. We obtained approval from our local institutional review board that operates in accordance with the Declaration of Helsinki. All participants signed informed consent forms. The trial was registered at clinicaltrials.gov (identifier: NCT01021254). The research was regularly reviewed by a data safety monitoring board.

Participants
Participants were recruited from outpatient psychiatry and general medicine clinics at UCSF and the outpatient psychiatry clinic at Kaiser Permanente in San Francisco and from the community. Clinics were located in large medical centers that serve an ethnically diverse population of privately and publicly insured adults in the San Francisco Bay Area. In addition to direct enrollment from these clinics, recruitment strategies included clinician referrals, flyers in public places, online advertisements, and local print advertisements in newspapers and on municipal buses and trains.

To be eligible for the study, participants had to be aged 18 years and older, fluent in English, meet DSM-IV diagnostic criteria for unipolar major depressive disorder based on an evaluation with the Structured Clinical Interview for DSM-IV diagnosis (SCID) [29] and have a score of 14 or greater on the 17-item Hamilton Depression Rating Scale (HAM-D17) [30]. Additionally, participants were required to be taking antidepressant medications with evidence of two or more adequate trials prescribed during the current episode assessed with the Antidepressant Treatment History Form (ATHF; described later) [33]. Participants were excluded for any of the following reasons: they (1) had a lifetime history of bipolar illness, schizophrenia or any psychotic disorder; (2) met DSM-IV criteria for substance abuse or dependence within 3 months prior to study entry; (3) were imminently suicidal, a danger to others or were currently exhibiting self-injurious behavior; (4) were in concurrent individual or group psychotherapy and not willing to discontinue these during the 8-week acute treatment duration of the study, or (5) had cognitive impairment, as defined by a score <25 on the Mini Mental State Examination [34]. Anxiety-related disorders and eating disorders were acceptable for enrollment if these cooccurring disorders were stable and not the primary disorder.

Randomization
Randomization was stratified by gender using computer-generated numbers. Participants were randomized by person after baseline assessment. These assessments occurred within 2 weeks of treatment initiation to ensure participants continued to meet the inclusion criteria for depression severity listed above. Participants and facilitators were blind to hypotheses and were informed of treatment assignments by the study coordinator who was not blinded or involved in outcome evaluation.

Study Treatment Interventions
Both treatments followed detailed manuals and were administered to small groups (6–12 participants) once per week for 8 consecutive weeks. Groups lasted 2 h and 15 min, and in addition to weekly attendance, participants were assigned 45 min of homework 6 days per week. Participants in both conditions were encouraged to continue their antidepressant treatment as prescribed by their outside provider and to report any medication changes weekly throughout the study.

Although treatment conditions were led by two different sets of facilitators, both sets were extensively trained in their respective techniques, having at least 3 years of experience in their respective components. All passed trial runs of their interventions and had strong beliefs in the value of their approach. All group sessions were audio recorded, and three recordings per group were randomly selected and reviewed by external and internal evaluators. Internally, HEP facilitator fidelity was evaluated at UCSF by a trained reviewer (E.G.) and externally HEP facilitator fidelity was evaluated by an investigator at the University of Wisconsin, Madison (Wis., USA), where HEP was developed [33]. Similarly, MBCT fidelity was assessed through an internal UCSF reviewer (S.J.E.) and externally by an unaffiliated psychiatrist trained in MBCT. Based on reviews and results of adherence and competency measures [34] (MBCT competency, HEP adherence and HEP competency scales are unpublished but available from the first author), corrective feedback was provided by the principal investigator when needed to group leaders during weekly supervision meetings with each team of interventionists.
Mean adherence ratings per item were 1.79 (SD = 0.06) out of 2 points maximum for MBCT using the MBCT adherence scale [34] and 1.82 (SD = 0.04) out of 2 points maximum for the HEP adherence scale. The mean competency ratings per item were 4.95 (SD = 0.02) out of 5 points on the MBCT competency scale and 4.90 (SD = 0.04) out of 5 points on the HEP competency scale. All scores were in the highly acceptable range.

MBCT was adapted from the original manual developed by Segal et al. [37], with modifications for TRD for the purpose of this study (published elsewhere) [38]. Using guided meditations and exercises from CBT, participants were taught skills to identify cognitive distortions and to disengage from depression-focused ruminate thinking patterns through the use of nonjudgmental and present-focused awareness. MBCT utilizes body scans, sitting meditations, 3-min breathing spaces and mindful movement as core techniques. Group discussion of how individuals are developing their meditation practice occurs in each session. Psychoeducation regarding depression is also an important component. Home practice consisted of mindfulness meditation, mindfulness of everyday life and mindful movement. With our TRD population, we modified the original protocol several ways [30, 39], such as shortening the length of meditations to a maximum of 30 min, emphasizing mindful movement, exploring barriers to practice and incorporating a focus on acceptance of emotional events.

The original HEP manual had been developed for use in a non-clinical sample as a control condition for mindfulness-based stress reduction, the forerunner of MBCT [35]. The HEP manual was adapted for the TRD population by linking the components to having potential beneficial effects on mood in class discussions. Course content focused on aerobic exercise, functional movement, music therapy and dietary education. Participants were asked to do specific exercises (e.g. walking/stretching/agility training), receive nutritional education and participate in musical activities such as drumming or song writing. They were told that these activities could reduce stress and promote overall health. Home practice consisted of activities such as exercises, stretching, monitoring food intake with a daily journal and musical activities. For both MBCT and HEP, daily diaries of home practice were completed online.

Medication Treatments

During the trial, patients continued to be seen for medication visits an average of once monthly. No prohibitions were made for medication changes during the trial. We encouraged reporting of any changes in dose or medications, and recorded these.

Assessment of Outcomes and Measures

Depression. Patients were evaluated with the HAM-D$_{17}$ at baseline (week 0), midway through treatment (week 4) and after treatment (week 8) as our main focus. We examined the primary outcome of change in depression severity (overall percentage change in the total score from baseline), as well as secondary outcomes of treatment response ($\geq$50% decrease from baseline scores) and remission (scores after treatment $\leq$7 on the HAM-D$_{17}$). As a more global measure of change in depression severity, we utilized the Clinical Global Impression instrument [40] to assess improvement and severity changes at similar time points as the HAM-D$_{17}$.

Mediators of Treatment Efficacy

The theoretical basis for mechanisms of change in MBCT has been posited to include improved emotional regulation [39] related to enhanced mindfulness [42], increased self-compassion [28], reduced rumination [43] and decreased experiential avoidance of dysphoric feelings [44].

We explored these potential mediators of treatment efficacy at weeks 4 and 8 using four instruments. Mindfulness was evaluated using the 39-item 5 Facet Mindfulness Questionnaire (FFMQ) [45], a self-report measure that assesses mindfulness across 5 domains: observing, describing, acting with awareness, accepting without judgment and nonreactivity to internal experience. We used the total score from the 26-item Self-Compassion Scale (SCS) [46] to evaluate the level of self-compassion across 6 subscales including: self-kindness, self-judgment, common humanity, isolation, mindfulness and over-identification. We measured rumination with the Ruminative Response Scale (RRS) [47], a 22-item self-report scale. The reliability of the scale has been repeatedly confirmed for patients with depression, and high scores appear to be associated with a vulnerability factor to depression and predict depression severity in clinically depressed populations. Finally, experiential avoidance, which has been associated with increased levels of depression [48] and is defined by attempts to avoid thoughts, feelings, memories, physical sensations or other internal experiences, was assessed using the 9-item Acceptance and Action Questionnaire (AAQ) [49]. Increased acceptance and decreased avoidance have been associated with decreased depression [44].

Moderators of Treatment Efficacy

Psychiatric factors that may impact depression severity and treatment resistance were analyzed as potential moderators of treatment. These included demographic factors such as ethnicity, minority and socioeconomic status as well as psychiatric history factors such as age at depression onset, number of lifetime depressive episodes and comorbid psychiatric illness (e.g. personality disorders [50]). We also examined other potential moderators of outcome such as anxiety utilizing the State and Trait Anxiety Inventory [51], current stress with the Perceived Stress Scale (PSS) [52] and early childhood stress with the Childhood Trauma Questionnaire (CTQ) [53]. The CTQ assesses physical, sexual or emotional abuse and trauma or neglect, all of which are widely reported to influence depression [54]. We also assessed the presence of personality disorders using the SCID-II [55]. Finally, since patients’ expectations and beliefs about treatment may affect outcome [56], we included a measure of treatment expectancy and credibility, the Treatment Rationale Scale [57], which participants completed after session 1. The scale includes expectancy of treatment success and how confident the individual would be in recommending the treatment to a friend.

Statistical Analysis

Standard summary statistics were used to describe the study sample. Treatment conditions were compared on the primary outcome measure (percent change in depression severity from baseline on the HAM-D$_{17}$). We regarded this continuous variable as the most sensitive indicator of change in this population. We also assessed the dichotomous variables of percent responding and percent achieving remission at the end of treatment (week 8). We used generalized logistic and linear regression models and generalized estimating equations to account for nonindependence due to experiencing a group-based intervention. Covariate terms were added to the models testing primary efficacy after
screening for their relationship to the outcomes. Mediational effects were tested using the methods outline by MacKinnon [58] and implemented in Mplus v7.2. The model of mediation of treatment condition on percent reduction in severity was estimated and tested using the potential mediators (FFMQ scales, SCs, RRS and AAQ) measured at week 4. As 4 weeks may not have been sufficient for more thorough skill acquisition, the model was reestimated using those measures when taken at the end of treatment. Moderation was tested by including interaction terms in the statistical models. Model estimation and testing was conducted under the intent-to-treat principle, and all available data were used in analyses. Power was estimated for each hypothesis and the methods of estimation were aligned with the methods of analysis. Estimates of effect size were derived from pilot work and the extant literature. Primary analyses were conducted using SAS v9.4.

Results

A total of 229 adults provided informed consent and were screened for eligibility [see fig. 1 (Consort diagram) for details]. Of the total screened, 173 subjects were eligible for randomization to either MBCT or HEP. Summary demographic statistics and clinical characteristics of the sample of the 173 participants by treatment condition at the baseline assessment are shown in table 1. There were no statistically significant differences between the conditions in any demographic or clinical variable. Table 1 shows that the sources of participant recruitment were distributed evenly across the three sites, and the treatment conditions did not differ by recruitment site \( \chi^2 = 0.33, p = 0.84 \). In the MBCT cohort, there were 4
changes in medication during the 8 weeks, while in the HEP cohort there were 5 changes.

In another secondary analysis, eight measures considered potential confounders in the preliminary plan were screened. Of the eight, four were significantly correlated with outcomes: PSS total score, presence of a personality disorder, and state and trait anxiety on the State-Trait Anxiety Inventory. The number of treatment sessions attended were inversely correlated with percent change in depression severity. These five measures were included as covariates in the main models testing treatment effect.

By the end of treatment, week-8 data were available on 83.7% of the HEP participants and 87.3% of the MBCT participants, indicating a dropout rate of 16.3% (HEP) and 12.7% (MBCT). The most common dropout reasons in both conditions, in descending frequency, were the following: (1) lost to contact; (2) worsened medical condition; (3) time burden; (4) scheduling conflict, and (5) change in family circumstance.

A multivariate analysis showed that relative to the HEP condition, the MBCT condition produced a significantly greater mean percent reduction in depression severity (36.6 vs. 25.3%; p = 0.01) at the endpoint. MBCT also produced a significantly greater number of treatment responders (30.3 vs. 15.3%; p = 0.03) based on the HAM-D 17 at the endpoint (table 2). Although also favoring MBCT, no statistically significant difference was found for rates of remission (22.4 vs. 13.9%; p = 0.15). Additionally, no differences based on gender were found.

In these models, state anxiety was related to diminished HAM-D 17 percent reduction in symptom severity and a lower response rate, but did not interact with condition. Stress (PSS score) and the presence of a personality disorder were related to diminished percent reduction, but did not interact with condition.

Similar modeling was used to test for effects of childhood trauma on the outcomes – both as predictors of outcome and moderators of the treatment effect. Only the emotional abuse subscale reached statistical significance in relationship to both response (p = 0.05) and remission (p = 0.03) with participants who responded and those who remitted having lower mean values. No evidence of moderation, as evidenced by a significant interaction term, was seen. Similar modeling was used to test for effects of ≥3 episodes of depression, but we did not

### Table 1. Demographic and clinical characteristics between treatment conditions of study participants with TRD at baseline (n = 173)

<table>
<thead>
<tr>
<th>Variable</th>
<th>MBCT</th>
<th>HEP</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Demographic characteristics</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Age, years</td>
<td>47.1 ± 13.46</td>
<td>45.2 ± 11.19</td>
</tr>
<tr>
<td>Range</td>
<td>21–69</td>
<td>16–85</td>
</tr>
<tr>
<td>Education, years</td>
<td>15.6 ± 2.01</td>
<td>15.9 ± 2.15</td>
</tr>
<tr>
<td>Female, %</td>
<td>75.9</td>
<td>76.7</td>
</tr>
<tr>
<td>Hispanic, %</td>
<td>8.3</td>
<td>11.6</td>
</tr>
<tr>
<td>Race, %</td>
<td>80.7</td>
<td>79.5</td>
</tr>
<tr>
<td>Asian, %</td>
<td>4.8</td>
<td>6.4</td>
</tr>
<tr>
<td>African American, %</td>
<td>8.4</td>
<td>7.7</td>
</tr>
<tr>
<td>Other, %</td>
<td>6.1</td>
<td>5.3</td>
</tr>
<tr>
<td>Employment, %</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Full time</td>
<td>28.7</td>
<td>39.5</td>
</tr>
<tr>
<td>Part time</td>
<td>18.4</td>
<td>22.1</td>
</tr>
<tr>
<td>Unemployed/student/homemaker</td>
<td>36.8</td>
<td>31.4</td>
</tr>
<tr>
<td>Retired, %</td>
<td>11.5</td>
<td>7.0</td>
</tr>
<tr>
<td>Disability, %</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mental health</td>
<td>19.5</td>
<td>15.1</td>
</tr>
<tr>
<td>Physical</td>
<td>8.1</td>
<td>7.0</td>
</tr>
<tr>
<td>Marital status, %</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Single, %</td>
<td>39.8</td>
<td>47.5</td>
</tr>
<tr>
<td>Married/cohabiting</td>
<td>31.3</td>
<td>34.6</td>
</tr>
<tr>
<td>Divorced/separated/widowed</td>
<td>28.9</td>
<td>17.9</td>
</tr>
<tr>
<td><strong>Clinical characteristics</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Age at depression onset, years</td>
<td>18.8 ± 10.86</td>
<td>21.7 ± 13.24</td>
</tr>
<tr>
<td>Depressive episodes, total n</td>
<td>3.8 ± 2.61</td>
<td>3.5 ± 2.37</td>
</tr>
<tr>
<td>Length of current depressive episode, months</td>
<td>84.4 ± 119.51</td>
<td>78.5 ± 93.5</td>
</tr>
<tr>
<td>HAM-D 17 score</td>
<td>18.3 ± 3.36</td>
<td>17.4 ± 3.48</td>
</tr>
<tr>
<td>Chronic depression (≥2 years of duration)</td>
<td>54.0</td>
<td>64.0</td>
</tr>
<tr>
<td>&lt;3 episodes, %</td>
<td>35.2</td>
<td>40.3</td>
</tr>
<tr>
<td>≥3 lifetime episodes, %</td>
<td>64.8</td>
<td>59.7</td>
</tr>
<tr>
<td>Previous depressive episodes, n</td>
<td>3.80</td>
<td>3.52</td>
</tr>
<tr>
<td>Adequate ATHF trials in episode, n</td>
<td>2.9</td>
<td>3.06</td>
</tr>
<tr>
<td>Previous psychiatric hospitalizations, %</td>
<td>16.1</td>
<td>18.6</td>
</tr>
<tr>
<td>History of suicide attempt, %</td>
<td>21.8</td>
<td>22.1</td>
</tr>
<tr>
<td>Comorbid anxiety disorder, %</td>
<td>56.3</td>
<td>63.9</td>
</tr>
<tr>
<td>Eating disorder, %</td>
<td>13.8</td>
<td>11.6</td>
</tr>
<tr>
<td>Personality disorder, %</td>
<td>11.5</td>
<td>16.2</td>
</tr>
<tr>
<td>Recruitment source, %</td>
<td></td>
<td></td>
</tr>
<tr>
<td>General internal medicine</td>
<td>34.5</td>
<td>36.1</td>
</tr>
<tr>
<td>Psychiatry clinic</td>
<td>43.7</td>
<td>39.5</td>
</tr>
<tr>
<td>Community</td>
<td>21.8</td>
<td>24.4</td>
</tr>
</tbody>
</table>

Means ± SD, numbers or percentages. There were no significant differences between conditions on any demographic or clinical variable.

### Table 2. Percent reduction, response rate and remission rate in HAM-D 17 by mid- and end-of-treatment assessment points

<table>
<thead>
<tr>
<th>Variable</th>
<th>Reduction, %</th>
<th>Response, %</th>
<th>Remission, %</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>W4</td>
<td>W8*</td>
<td>W4</td>
</tr>
<tr>
<td>HEP</td>
<td>23</td>
<td>25.3</td>
<td>10.1</td>
</tr>
<tr>
<td>MBCT</td>
<td>24</td>
<td>36.6</td>
<td>10.1</td>
</tr>
</tbody>
</table>

W = Week. * p < 0.05.
find evidence of a relationship of this characteristic to outcome.

Treatment expectancy did not differ statistically between conditions after one session. The MBCT group and HEP group did not differ in the two treatment expectancy items: (1) how likely the treatment would be successful \((p = 0.62)\) and (2) how confident the individual would be in recommending the treatment to a friend \((p = 0.23)\).

In terms of other factors that could have affected outcomes, we found no statistically significant differences in medication changes (e.g. number of increased or decreased dosages, or medication switches or augmentations) between conditions or evidence that such changes were related to clinical outcomes. Moreover, the number of ATHF adequate prior treatment trials did not significantly differ between the two conditions [MBCT: mean 2.90 (SD 1.31) and HEP: mean 3.08 (SD 1.37)], indicating similar levels of treatment resistance. Finally, although all of the putative mediators (table 3) changed in the direction anticipated (i.e. increased mindfulness, reduced rumination, enhanced self-compassion and decreased experiential avoidance), only the observing subscale of the FFMQ showed a greater change with MBCT than HEP \((p = 0.019)\). Moreover, the mediation models did not indicate any of these variables drove the observed clinical change in either arm of the study.

We also assessed a number of secondary variables over a 52-week follow-up interval. The Clinical Global Impression demonstrated beneficial effects for both conditions, but MBCT had more favorable changes for improvement \((p = 0.001)\) and decreased severity \((p = 0.038)\) at 8 weeks. By 52 weeks, there was no significant difference between conditions, although both continued to be significantly improved from baseline. The presence of increased perceived stress, personality disorder and short-term or trait anxiety predicted decreased response for both conditions. There were no significant differences between conditions at 24, 36 and 52 weeks for the putative mediators of avoidance, rumination, mindfulness or self-compassion, although all continued to be significantly improved from baseline. Clinical outcomes at 52 weeks using the HAM-D_{17} indicated that MBCT was significantly greater for percent responding but not for percent reduction or remitting. These findings are illustrated in figure 2.

**Discussion**

This study is the first full-scale RCT to investigate MBCT as an adjunctive treatment with medications for individuals with TRD who are currently depressed. In the current study, using a manualized and previously published control condition, HEP [35, 59], both arms demonstrated high rates of compliance and low dropout rates for a TRD population [60]. MBCT produced a significantly greater percent decrease in depression severity and response rate than HEP at the 8-week endpoint. Although the remission rate was numerically greater for MBCT than HEP, the difference was not statistically significant. In this treatment-resistant population comprised of patients who had failed to remit despite an average of three medication trials prior to study entry, low remission rates would be expected and differences between low rates would be difficult to detect in a sample of this size. In addition, although our entry criteria were a HAM-D_{17} score \(\geq 14\) (the same cutoff used in the STAR*D study), for both MBCT and HEP mean HAM-D_{17} entry scores

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**Table 3. Putative mediators over time**

<table>
<thead>
<tr>
<th>Putative mediators</th>
<th>Baseline</th>
<th>4 weeks</th>
<th>8 weeks</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>MBCT</td>
<td>HEP</td>
<td>MBCT</td>
</tr>
<tr>
<td></td>
<td>SD</td>
<td>SD</td>
<td>SD</td>
</tr>
<tr>
<td>AAQ total score</td>
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<tr>
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were >17, indicating the population was significantly depressed upon study initiation.

To place our results in context, we can compare our findings to other psychotherapeutic efforts in depression and TRD. For example, our dropout rates of 16.3% (HEP) and 12.7% (MBCT) compared favorably to the rate of 24.3% for CBT in unipolar depressed outpatients in a recent meta-analysis [61]. Brakemeier et al. [62] completed an uncontrolled study of CBASP in inpatients suffering chronic depression and TRD. Our results were modest in comparison. Our 8-week response/remission rates of 30.7/22.4% compared to the rates of 75.7/40.0% of Brakemeier et al. [62] at 12 weeks. However, their inpatient treatment offered a longer and more intense milieu treatment. Our 52-week responder rate (43.7%) was similar to their 48% rate. Perhaps our results can be better assessed and compared to other randomized outpatient trials of TRD management. Our 8-week results demonstrate outcomes that were superior to the 12-week medication steps 3 or 4 of the STAR*D study. That study found response/remission rates of 24.0/14.1% at step 3 and 23.2/13.1% at step 4 [63]. In a trial of CBASP with patients having chronic depression and TRD, rates for response/remission were 22.5/15.0% [12], whereas in our study MBCT produced rates that were again superior at 8 weeks. Taken as a whole, these studies suggest that expectations must be tempered by what is feasible, yet there is room for cautious optimism in undertaking treatment in TRD. Moreover, recently, Kuyken et al. [64] demonstrated that MBCT with discontinuation of antidepressants was not inferior to maintenance medications in terms of preventing relapse. Our study raises the possibility that MBCT offers improvement similar to additional medication trials for those with TRD with a potential for relapse prevention as well.

This study strengthens the finding of a previous study by Chiesa et al. [65] that had similar main effects of MBCT, but was based on a smaller, preliminary sample with less stringent requirement for TRD that utilized a psychoeducational control. In a review of treatment approaches for TRD, Carvalho et al. [66] noted that cognitive therapy appears to be an effective adjunctive treatment to medication for TRD. Our report broadens consideration of psychotherapeutic approaches for this challenging condition to include MBCT.

Identifying appropriate controls for mindfulness interventions has been challenging and researchers have often utilized relatively weak controls such as wait lists. We chose HEP as the comparator condition for several reasons. First, the University of Wisconsin and NCCAM (National Center for Complementary and Alternative Medicine) specifically developed it to be a credible, structurally equivalent control for mindfulness intervention such as mindfulness-based stress reduction and its derivative, MBCT. Second, it provides the nonspecific elements of therapy, including group support, attention (time spent with a therapist), home practice, diminished stigma and provision of hope. Third, exercise, the central component of the HEP, was initially conceived as a control for the movement component of mindfulness interventions (e.g. mindful walking and yoga) and as a way of building a credible alternative condition. Since exercise has been shown to have significant antidepressant effects [67], this may help to explain why the HEP group improved. The fact that HEP demonstrated some change in depression levels after 8 weeks could have been expected due to its nonspecific effects. In addition, HEP has been shown to be as efficacious as mindfulness-based stress reduction for reducing perceived stress, a factor contributing to depression, in two previous studies [35, 59]. In our

Fig. 2. HAM-D17 outcomes over a 52-week follow-up.
prettrial evaluations, we observed that HEP has a behav-
ioral activation element that was not described in its ini-
tial research. Encouraging participants to engage in ac-
tivities such as exercise, listening to music or eating nu-
ritionally could all have helped make HEP a credible in-
tervention as well as potentially having beneficial mood
effects. Finally, structural elements of the design of HEP
also made it a strong comparator condition. For example,
HEP was equivalent to MBCT in number of sessions and
home practice; it was also led by two highly motivated fa-
cilitators (>3 years teaching meditation or the compo-
ents of HEP) for each group. Although the same inter-
ventionists did not deliver both conditions, both MBCT
and HEP interventionists believed their procedures could
help alleviate depression.

How MBCT may achieve its effects needs further elu-
cidation [68]. While the putative mediators we examined
changed in the expected direction with MBCT, none of
them met the statistical test of mediation. Although it is
possible that we lacked sufficient power to detect a me-
diator in a secondary analysis, the lack of a significant ef-
fact of our hypothesized mediators raises the possibility
that another factor may be a key ingredient. There could
be another variable like attentional control, autobi-
ographical memory or self-discrepancy that may have
played a role, but we faced a limit in both respondent bur-
den and the ability to usefully analyze a larger number of
variables.

Without being clear on what the actual mediator was,
we cannot be sure why we did not detect it. We did exam-
ine the most commonly reported mediators of mind-
fulness, self-compassion, rumination and experiential
avoidance, and failed to identify any as passing the test of
mediation [56]. Moreover, it is possible that the current
measures of mindfulness may not accurately capture
what actually changes with treatment. For example, there
is evidence that exercise, such as included in HEP, may
have an impact on what we measure as mindfulness.
Mothes et al. [69] demonstrated that regular aerobic ex-
ercise can increase measures of dispositional mindfulness
using the Mindfulness Attention Awareness Scale [68].
Such findings point to the importance of refining our def-
initions of mindfulness and being aware of how well they
match the sample being investigated [69], i.e. are mea-
sures equally valid in community, and clinical and medi-
tator samples.

In addition to the above factors, the long-term anti-de-
pressant treatment that our TRD population underwent
may have washed out a clearer picture of the effects of our
putative mediators. As Fava [72] points out, certain fea-
tures of long-term antidepressant treatment may play a
role in contributing to treatment resistance that may in-
terfere with psychotherapeutic responsiveness. These fea-
tures could potentially influence the effects of mediators,
making it more difficult to identify which ones may be
playing a significant role.

We examined a number of potential predictors of the
outcomes and found that the following factors did not
have a significant effect on outcomes: duration of depres-
sive episodes, number of episodes, presence of disability,
early onset of illness, sociodemographic group, education
level, medical illness, depression severity, amount of
homework completed or recruitment site. Of particular
note, some studies of MBCT for relapse prevention [19,
29] have found efficacy only in individuals with three or
more episodes of depression. We did not find such a re-
lationship, perhaps because the duration of episodes in
our sample produced a picture of chronicity that shared
features of those with three or more episodes.

While the above factors did not appear to have a sig-
ificant impact on outcome, several other factors were
predictors of outcome independent of treatment. Data
showed that a greater number of sessions attended were
positively correlated with a greater treatment effect. We
also found that the presence of a personality disorder or
anxiety on both a state and trait level influenced outcome
adversely. This is consistent with other studies showing
the impact of personality disorders and anxiety on de-
pression treatment [73]. Perceived stress also had a sig-
nificant adverse effect. Stress has been regarded as a key
contributor to depression, so this effect is not unantici-
pated. None of the above factors showed an interaction
with treatment condition. Finally, one subscale of the
CTQ, emotional abuse, also showed a relationship to ad-
verse outcome but not differentially by treatment condi-
tion. This was in contrast to one study suggesting MBCT
was more efficacious than a psychoeducational control
condition when childhood trauma was present [74].

The long-term outcomes indicate that the beneficial
effects of MBCT occurred early and were maintained over
the 52 weeks, while HEP effects appeared to have a later
onset. In both conditions, people got better after 8 weeks
with sustained improvement over 1 year, although the
improvements were more substantial in MBCT at the
8-week time point, the primary focus of this study. While
MBCT remained numerically superior for percent reduc-
tion in HAM-D17 and percent responders, there was no
statistical difference at 52 weeks. The decrease in con-
dition separation over 52 weeks is not atypical for psycho-
therapy studies, which often report that benefits are
maintained over time but that condition differences do not continue after intervention cessation [75, 76]. Although powered for 8-week differences in clinical outcomes (i.e. percent reduction in HAM-D17), there is a possibility that with time, attrition effects occurred with participants who were not doing as well and may have dropped out so that only individuals who were doing better agreed to complete long-term follow-up assessments in both conditions; such a process may have decreased the separation. Additionally, although HEP was not chosen as a therapeutic control, it has components that could have produced beneficial antidepressant effects, e.g. exercise has been identified as improving depressive disorders, particularly as an augmentation intervention [77], and these effects may also have played a role in reducing the separation of the conditions by 52 weeks.

The study had a number of strengths. It utilized two manualized interventions, MBCT and HEP, that offered structural equivalence. HEP provided a credible, active comparator for MBCT as evidenced by the week 1 Therapeutic Rationale Scale scores for each condition. We carefully monitored facilitator fidelity to each model and competency. We note, however, that only the MBCT adherence scale [36] was published, and the measures of MBCT competency and HEP adherence and competency have not been validated or published thus far. However, on most of the items rated by two people, there was perfect agreement on a 5-point scale, and when they disagreed, they differed by only 1 point. We should also comment that the principal investigator, trained in both MBCT and HEP, delivered feedback to both MBCT and HEP interventionists. Since he knew the hypothesis, this introduced a potential for bias. Another limitation of the study was that the two interventions were delivered by two different sets of individuals. However, it would be difficult to find one set of individuals who would be skilled across both domains.

Another limitation relates to the 8-week length of treatment. It is possible that with a longer intervention, or mandated monthly alumni meetings, long-term results may have improved. Longer treatment lengths may have been particularly useful for individuals with chronic depression.

We screened participants to meet rigorous criteria for TRD, and they represented a seriously ill population. The participants, although blind to study hypotheses, were not masked to intervention assignment that could have introduced some bias. The study title ‘Practicing Alternative Techniques to Heal Depression’ was chosen to prevent introducing a perceived investigator bias. Another factor that may have played a role in outcome was our decision not to limit medication changes during the treatment period. Medication changes were tracked on a weekly basis and we found no differences between conditions on the number of medication changes that occurred. Furthermore, there was no evidence of a relationship between medication changes and treatment outcomes.

Patient preference for treatment assignment was not measured before randomization, and this may have played a role in outcome. Some studies, however, suggest that treatment preferences in RCTs may not significantly affect outcome [78]. In addition, Steidtmann et al. [79] indicate that patient preference may act as a moderator, but in unexpected ways. For example, in their study, individuals expressing a preference for medication treatment of depression and then received such treatment had higher attrition rates than those who had not expressed such a preference receiving that treatment. Individuals’ past psychotherapy and medication treatment histories and illness belief models are additional factors that affect patient preference. Hence, although the relationship between patient preference and outcome is complex, it would have been useful to have collected initial preference prior to randomization in order to assess potential effects.

In summary, MBCT appears to have specific effects that make it a useful treatment for patients with TRD. Many depressed individuals are interested in learning nonpharmacological and empowering approaches to control their illness. Moreover, in contrast to the suggestion that actively depressed patients could not concentrate well enough to utilize MBCT [37], our findings did not support this belief. Subsequent papers will describe the changes in brain function associated with MBCT and HEP. The suffering related to TRD is pervasive, long lasting and challenging. New approaches such as MBCT for augmentation of medication management may play an important role in the therapeutic armamentarium.

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**Disclosure Statement**

All authors report no conflict of interest.
References